Fitting a geometric graph to a protein–protein interaction network

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**Abstract**

**Motivation:** Finding a good network null model for protein–protein interaction (PPI) networks is a fundamental issue. Such a model would provide insights into the interplay between network structure and biological function as well as into evolution. Also, network (graph) models are used to guide biological experiments and discover new biological features. It has been proposed that geometric random graphs are a good model for PPI networks. In a geometric random graph, nodes correspond to uniformly randomly distributed points in a metric space and edges (links) exist between pairs of nodes for which the corresponding points in the metric space are close enough according to some distance norm. Computational experiments have revealed close matches between key topological properties of PPI networks and geometric random graph models. In this work, we push the comparison further by exploiting the fact that the geometric property can be tested for directly. To this end, we develop an algorithm that takes PPI interaction data and embeds proteins into a low-dimensional Euclidean space, under the premise that connectivity information corresponds to Euclidean proximity, as in geometric-random graphs. We judge the sensitivity and specificity of the fit by computing the area under the Receiver Operator Characteristic (ROC) curve. The network embedding algorithm is based on multi-dimensional scaling, with the square root of the path length in a network playing the role of the Euclidean distance in the Euclidean space. The algorithm exploits sparsity for computational efficiency, and requires only a few sparse matrix multiplications, giving a complexity of \(O(N^2)\) where \(N\) is the number of proteins.

**Results:** The algorithm has been verified in the sense that it successfully rediscovers the geometric structure in artificially constructed geometric networks, even when noise is added by re-wiring some links. Applying the algorithm to 19 publicly available PPI networks of various organisms indicated that: (a) geometric effects are present and (b) two-dimensional Euclidean space is generally as effective as higher dimensional Euclidean space for explaining the connectivity. Testing on a high-confidence yeast data set produced a very strong indication of geometric structure (area under the ROC curve of 0.89), with this network being essentially indistinguishable from a noisy geometric network. Overall, the results add support to the hypothesis that PPI networks have a geometric structure.

**Availability:** MATLAB code implementing the algorithm is available upon request.

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**1 Introduction**

Large, complex networks (also called graphs) arise in a vast array of applications (Newman, 2003). Efforts to develop models that describe and summarize complex networks have focused on various network features such as motifs (Milo et al., 2002), graphlets (Pržulj et al., 2004) and graphlet degree distributions Pržulj (2006), clustering coefficients (Watts and Strogatz, 1998), pathlengths (Watts and Strogatz, 1998) and degree distributions (Khanin and Wit, 2006; Newman, 2003; Thomas et al., 2003).

Studying protein–protein interaction (PPI) networks has recently become possible due to advances in experimental high-throughput technologies such as yeast-2-hybrid (Y2H) (Y2H) (Ito et al., 2000; Uetz et al., 2000), tandem affinity purification (TAP) (Gavin et al., 2002) and high-throughput mass spectrometric protein complex identification (HMS-PCI) (Ho et al., 2002). A significant amount of experimental PPI network data for several organisms has already been generated. (Gavin et al., 2002; Giot et al., 2003; Ho et al., 2002; Ito et al., 2000; Krogan et al., 2006; Li et al., 2004; Rual et al., 2005; Stelzl et al., 2005; Uetz et al., 2000).

Understanding the patterns of intricate wiring in PPI networks is clearly of great importance for basic biological understanding, and also has the potential to feed back into the strategies for optimal interactome detection (Lappe and Holm, 2004). Further benefits of an accurate PPI model include (a) generation of synthetic datasets of any size in order to test computational algorithms, (b) detection of false positives and false negatives, (c) possible insights into the evolutionary processes that created the network and (d) convenience of representing complex networks in terms of a small number of model parameters and thereby distinguishing between networks for different organisms. Thus, modelling of PPI networks has become an active research area and several different random graph models have already been suggested (Grindrod, 2002; Grindrod and Kibble, 2004; Morrison et al., 2006; Pržulj and Higham, 2006; Pržulj et al., 2004; Thomas et al., 2003; Vazquez et al., 2001). Among them are geometric random graphs (Pržulj et al., 2004, 2006) in which nodes correspond to uniformly randomly distributed points in a low-dimensional Euclidean space and edges exist between pairs of nodes in the graph if the corresponding points in the space are close enough (within some radius \(\epsilon\)) according to the Euclidean distance norm. Other models include: Erdős–Rényi random graphs (Erdős and Rényi, 1959, 1960), generalized random graphs...
(Bender and Canfield, 1978), small-world (Watts and Strogatz, 1998), scale-free (Barabási and Albert, 1999; Simon, 1955) and stickiness (Pržulj and Higham, 2006) networks.

Our research focuses on geometric random graphs. The key observation that drives our work is that, in contrast with other putative PPI models, the geometric structure can be examined constructively. To test whether a given PPI network has a geometric structure, rather than measuring local and global statistics of the PPI network and comparing these with local and global statistics of random geometric graphs, a more direct question can be addressed:

Can we represent the given PPI network as a geometric graph by embedding the proteins in $\mathbb{R}^2$, $\mathbb{R}^3$ or $\mathbb{R}^4$ and finding an $\epsilon$ such that proteins are connected if and only if they are $\epsilon$-close?

There are two main themes to this work. The first theme is designing and testing an algorithm to discover whether a given PPI network has an underlying geometric structure. This theme is dealt with in Sections 2–4. The second theme, covered in Section 5, is to use this tool to study PPI networks.

We remark that the reverse engineering problem considered here is related to the general, but less well-defined, tasks of ordering and clustering (Grindrod, 2002; Grindrod and Kibble, 2004; Titz et al., 2004). Spectral (eigenvalue/eigenvector-based) algorithms have proved successful for ordering and clustering (Grindrod and Kibble, 2004; Higham, 2003), and this provides motivation for a spectral algorithm to address the geometric embedding issue.

## 2 THEORETICAL BASIS

As in previous studies (Pržulj, 2006; Pržulj et al., 2004), we focus on non-periodic, uniform, Euclidean geometric random graphs in $\mathbb{R}^2$, $\mathbb{R}^3$ and $\mathbb{R}^4$. These are defined as follows [see Penrose (2003) for further details about geometric random graphs and their properties]. In the two-dimensional (2D) case, to create a network (an undirected, unweighted graph) with $N$ nodes we place a set of $N$ points, $\{x^{(i)}\}_{i=1}^{N}$, uniformly in the unit square; that is, each $x^{(i)} \in \mathbb{R}^2$ has its two components drawn independently from the uniform (0,1) distribution, and all points are generated independently. Then, for each pair of points $(x^{(i)}, x^{(j)})$, we create an edge between nodes $i$ and $j$ of the geometric random graph if and only if $|x^{(i)} - x^{(j)}|_2 \leq \epsilon$, where $|x|_2$ denotes Euclidean distance and $\epsilon > 0$ is a parameter. In other words, nodes $i$ and $j$ are linked if and only if points $i$ and $j$ are within Euclidean distance $\epsilon$. The process is illustrated in the upper left picture of Figure 1. The three-dimensional (3D) and four-dimensional (4D) cases are defined analogously, by placing points in $\mathbb{R}^3$ and $\mathbb{R}^4$.

Our algorithm makes use of ideas from multi-dimensional scaling (MDS). We summarize here the necessary details, referring the reader to (Cox and Cox, 1994) for further information and historical references, and (Kaski et al., 2003; Taguchi and Oono, 2004) for examples where MDS has been used in bioinformatics.

Suppose that, for a set of $N$ objects, we are given the pairwise Euclidean distances $d_{ij}$ between all pairs, and we are asked to find a set of $N$ vectors $\{x^{(i)}\}_{i=1}^{N}$ in $\mathbb{R}^m$ such that

$$\|x^{(i)} - x^{(j)}\|_2 = d_{ij}, \text{ for all } i,j.$$  \hspace{1cm} (1)

In other words, our task is to find locations in $\mathbb{R}^m$ for the objects so that the pairwise distances are respected. Finding the smallest dimension $m$ for which a solution is possible may be regarded as part of the problem. In our context, we will think of the dimension as being fixed at 2, 3 or 4, and we will be seeking $N$ locations $\{x^{(i)}\}_{i=1}^{N}$ for which the constraints (1) are well approximated in a sense that will be made precise.

Given data $\{d_{ij}\}$ that respects the triangle inequality, double centering produces the symmetric, positive semi-definite matrix $A \in \mathbb{R}^{N \times N}$ defined by

$$a_{ij} = -\frac{1}{2} \left( d_{ij}^2 - \frac{1}{n} \sum_{k=1}^{n} d_{ik}^2 - \frac{1}{n} \sum_{k=1}^{n} d_{kj}^2 + \frac{1}{n^2} \sum_{k=1}^{n} \sum_{l=1}^{n} d_{kl}^2 \right).$$  \hspace{1cm} (2)

Letting $X \in \mathbb{R}^{m \times N}$ be the matrix whose $j$th column is $x^{(j)}$, it may then be shown that

$$X^T X A = A \Rightarrow \|x^{(i)} - x^{(j)}\|_2 = d_{ij}, \text{ for all } i,j.$$  \hspace{1cm} (3)

Now $A$ has the real Schur decomposition (Golub and Van Loan, 1996) $A = U^T \Sigma U$, where $U \in \mathbb{R}^{N \times N}$ is orthogonal (its rows are eigenvectors of $A$) and $\Sigma = \text{diag}(\sigma_i)$ is diagonal with diagonal entries ordered high-to-low (these are the eigenvalues of $A$). We then see that a solution $X$ in (3) may be computed as

$$X = \Sigma^{1/2} U.$$  \hspace{1cm} (4)
To find an ‘optimal’ approximation such that $x^{(i)} \in \mathbb{R}^r$, we may truncate (4) using only the $r$ most positive eigenvalues, so that

$$
\hat{X} = \begin{bmatrix}
\sqrt{\sigma_1 u_{ij}^{(k)}} & \cdots & \cdots \\
\vdots & \ddots & \cdots \\
\sqrt{\sigma_r u_{ij}^{(k)}} & \cdots & \cdots 
\end{bmatrix},
$$

(5)

where $u_{ij}^{(k)} \in \mathbb{R}^n$ is the $k$th row of $U$. This is optimal in the sense that $\hat{X}$ is the closest rank-$r$ matrix to $X$, in any orthogonally invariant norm (Golub and Van Loan, 1996).

3 INITIAL ALGORITHM

In this section, we outline the main ideas behind the algorithm and show how we propose to evaluate its accuracy. Our task is to embed proteins into $\mathbb{R}^2$, $\mathbb{R}^3$ or $\mathbb{R}^4$ given the PPI network. Rather than Euclidean distances, we have only $\{0,1\}$ connectivity information. For this reason, we will use a function of the pathlength in lieu of Euclidean distance. By construction, if two nodes are connected (pathlength $1$) then their Euclidean distance is between zero and $\epsilon$. Similarly, a pathlength of two indicates that an intermediate node, $C$, is $\epsilon$-close to both $A$ and $B$, with the Euclidean distance between $A$ and $B$ lying somewhere between $\epsilon$ and $2\epsilon$. In the absence of exact distance information we will adopt the heuristic that a ‘typical’ configuration has a right angle for the angle $ABC$, and assume that a typical length-two path corresponds to a distance of $\sqrt{2}\epsilon$. The square root can also be regarded as an attempt to compromise between the opposing factors where (1) one of the distances $A$-to-$C$ or $C$-to-$B$ is much less than $\epsilon$ and (2) the nodes $A$, $B$ and $C$ are co-linear. More generally, we will use the square root of the graph pathlength in lieu of Euclidean distance, so that $d_{ij}$ in (1) is taken to be $\sqrt{\text{path}_{ij}}$, where $\text{path}_{ij}$ denotes the pathlength between nodes $i$ and $j$. In practice, we tried several alternative monotonically increasing functions of $\text{path}_{ij}$ and found that the resulting algorithm was insensitive to this detail.

A minor issue is the natural scale-invariance of the problem—re-scaling $\epsilon$ and the distances $d_{ij}$ does not change the network. Because the traditional geometric model assumes that all points lie in the unit disk/cube, we will normalize the coordinate vectors in (5) so that

$$
\sqrt{\sigma_k u_{ij}^{(k)}} \rightarrow \frac{\sqrt{\sigma_k u_{ij}^{(k)}} - \min_k (\sqrt{\sigma_k u_{ij}^{(k)}})}{\max_k (\sqrt{\sigma_k u_{ij}^{(k)}}) - \min_k (\sqrt{\sigma_k u_{ij}^{(k)}})}.
$$

We now illustrate these ideas on data arising from a small geometric graph in $\mathbb{R}^2$ for which the results can be easily visualized. Here we took $N = 100$ nodes with coordinates drawn independently from the uniform $(0,1)$ distribution and joined nodes that were within Euclidean distance $\epsilon = 0.25$. The resulting graph is shown in the upper left picture of Figure 1. The six most positive eigenvalues of $A$ in (2), using $\sqrt{\text{path}_{ij}}$ for $d_{ij}$, were found to be $39.9, 31.3, 8.7, 7.4, 6.1$ and $4.0$. If we had used exact Euclidean distance information then, since we started with a geometric graph in $\mathbb{R}^2$, it would be possible to embed exactly in $\mathbb{R}^2$, so that only the first two eigenvalues would be non-zero. In using pathlength to approximate Euclidean distance, we have lost this property, but it is reassuring that the first two eigenvalues remain strongly dominant. The 2D embedding from the algorithm is shown in the upper right, and the lower pictures display the false positives and false negatives arising when a geometric graph with radius $\epsilon = 0.25$ is formed.

To measure the ability of the algorithm to recover the original network, we present a receiver operator characteristic (ROC) curve (Bradley, 1997; Tape, 2000) in Figure 2, marked Geometric/MDS. Here, we increased $\epsilon$ from 0 to $\sqrt{2}$ in small increments, and for each $\epsilon$ we generated the geometric graph arising from the MDS node placement as in the upper right picture of Figure 1. The horizontal axis is then defined as $1 - \text{specificity}$, that is, $1 - \text{TN}/(\text{TN} + \text{FP})$, and the vertical axis is defined as sensitivity, that is $\text{TP}/(\text{TP} + \text{FN})$. Here TN denotes the number of true negatives that is, the number of distinct pairs $i$ and $j$ for which there is no edge in the reverse engineered graph and there is no edge in the original graph. Similarly, TP, FN and FP denote the number of true positives, false positives and false negatives, respectively. With $\epsilon = 0$, we place no edges in the network, and hence we have perfect specificity, $1 - \text{TN}/(\text{TN} + \text{FP}) = 0$, but the worst possible sensitivity, $\text{TP}/(\text{TP} + \text{FN}) = 0$. The other extreme $\epsilon = \sqrt{2}$ connects all nodes, giving the worst possible specificity, $1 - \text{TN}/(\text{TN} + \text{FP}) = 1$, but perfect sensitivity, $\text{TP}/(\text{TP} + \text{FN}) = 1$. Increasing $\epsilon$ always improves sensitivity at the expense of specificity. Good performance corresponds to having a curve that rises rapidly, containing points close to $x = 0$, $y = 1$, and the area under the curve (AUC) is a widely-used measure of quality (Bradley, 1997; Tape, 2000). In Figure 2 we have an AUC of 0.988 for MDS curve.

For comparison, we have added two more ROC curves in Figure 2. First, for the same network, we show the effect of ‘randomly guessing’ links. Here, we take a biased coin that lands heads with probability $p$. For each pair of nodes we flip the coin and predict a link if the coin lands heads. As $p$ is varied this leads to the ROC curve labelled Geometric/Coin Flip, which has an AUC of 0.47. Next, we generated a network with
100 nodes that did not have an underlying geometric structure. Instead we used an Erdős-Rényi model (Erdős and Rényi, 1959, 1960) as in (Pržulj et al., 2004) where, for each pair of nodes, a link was inserted with independent probability 0.8. Applying our MDS-based algorithm produced the ROC curve labelled Erdos/MDS, which has an AUC of 0.64.

4 PRACTICAL ALGORITHM

For PPI networks, where the number of nodes is typically in the thousands, we propose setting

\[ d_{ij} = \begin{cases} \sqrt{\text{path}_{ij}}, & \text{if } \text{path}_{ij} \leq K, \\ K_{\text{max}}, & \text{otherwise}. \end{cases} \]  

(7)

where \( K \) and \( K_{\text{max}} \) are parameters in the algorithm. Here we have introduced a cutoff \( K \) with longer pathlengths rounded to a single value \( K_{\text{max}} \). Using the cutoff \( K \) in (7), rather than computing and recording the pathlength for all pairs, has three main advantages.

(1) By choosing \( K \) relatively small, the resulting algorithm can exploit sparsity in the original network. This is explained further below.

(2) The case where the network consists of two or more disconnected components (i.e. some pathlengths are infinite) is conveniently handled.

(3) The cutoff reflects the fact that for a true geometric graph in the unit cube, there is an upper bound on the maximum Euclidean distance.

We remark that, intuitively, it is clear that accurate information concerning near-neighbours is more important than information concerning distant nodes. In our experiments, we found that the results from the algorithm were not sensitive to the choice of \( K \) and \( K_{\text{max}} \), although, as explained below, the computational benefits can be dramatic.

Repeating the experiment in Figure 1, but with parameters \( K = 4 \) and \( K_{\text{max}} = 5 \), gave rise to leading eigenvalues 38.3, 30.1, \( \sqrt{10} \), \( \sqrt{9} \), and \( \sqrt{8} \), and an area under the ROC curve of 0.984, showing that retaining level-four path information is adequate in this case. Similar effects were observed more generally, so we used these values in all computations.

We now explain how our distance measure (7) allows sparsity to be exploited. To measure complexity, we assume that the number of connections per protein is fixed, independently of \( N \), and we consider asymptotics as \( N \to \infty \). We assume that subspace iteration (Golub and Van Loan, 1996) is used to compute eigenpairs of the matrix \( A \) in (2). The significant computational cost is then the formulation of a matrix-vector product; given some \( v \in \mathbb{R}^N \) compute the product \( Av \).

By construction the elements \( d_{ij}^2 \) in (2) now come from a matrix of the form \( P + K_{\text{max}} I^T \), where \( I \in \mathbb{R}^N \) denotes the vector of 1’s, and hence \( I^T I \in \mathbb{R}^{N \times N} \) is the matrix of 1’s, and \( P \) has values

\[ p_{ij} = \begin{cases} \sqrt{\text{path}_{ij}} - K_{\text{max}}, & \text{if } \text{path}_{ij} \leq K, \\ 0, & \text{otherwise}. \end{cases} \]  

(8)

Hence, \( p_{ij} \) is non-zero only when there is a path of length \( \leq K \) between proteins \( i \) and \( j \). In other words, \( P \) has the same sparsity pattern as the \( K \)th power of the adjacency matrix for the network. For a fixed value of \( K \), this means that \( P \) has the same sparsity as the original network. It then follows that the product \( Av \) may be written

\[ Av = -\alpha \sum_{j=1}^N (v_j)p_{ij} + \left( \frac{\alpha}{N} \right) v^T p \]  

(9)

where \( p_{ij} \in \mathbb{R}^N \) has \( (p_{ij})_i = \sum_{j=1}^N p_{ij} \). \( v \in \mathbb{R} \) denotes the average value \( v^T 1/N \) and \( \alpha \in \mathbb{R} \) denotes \( \sum_{j=1}^N \sum_{j=1}^N p_{ij} \). It follows that each step of the subspace iteration involves a matrix–vector multiply with a sparse matrix. In our computations, we stopped the iteration when successive eigenvector approximations agreed to within \( 10^{-3} \) in Euclidean norm.

Overall, given a protein–protein interaction network, our algorithm may be summarized as:

(1) Compute the pathlengths up to length \( K \).

(2) Compute the first two, three or four most positive

eigenvalues of \( A \), and the corresponding eigenvectors.

(3) Embed the nodes in \( \mathbb{R}^2 \), \( \mathbb{R}^3 \) or \( \mathbb{R}^4 \) using (5) and (6).

(4) Examine the accuracy of the embedding as \( \epsilon \) is varied.

To measure computational cost we note that a sparse matrix–vector multiply \( Av \) is an \( O(N) \) process. Step 1 of the algorithm can be achieved by forming the matrices \( A, A^2, A^3, \ldots, A^K \), which has, at most, an \( O(N^2) \) operation count. Step 2 costs \( O(N) \) per iteration of the subspace iteration algorithm. Step 3 is negligible. In Step 4, having generated the protein locations, computing the pairwise distance data \( \|[x^T] - [x^T]p_{ij} \|_{2} \) so that choices for \( \epsilon \) can be tested, is an \( O(N^2) \) task.

5 DATA AND RESULTS

5.1 PPI networks

Using high-throughput techniques such as Y2H (Ito et al., 2000; Uetz et al., 2000), TAP (Gavin et al., 2002) and HMS-PCI (Ho et al., 2002), a significant amount of experimental PPI data has been generated. Our algorithm has been applied to 19 PPI networks of four eucaryotic organisms: yeast *Saccharomyces cerevisiae*, fruitfly *Drosophila melanogaster*, nematode worm *Caenorhabditis elegans* and human. We used nine yeast, one fruitfly, three worm and six human PPI networks obtained from different studies that used different PPI detection techniques, as well as from curated databases (described below).

The high-confidence part of the yeast PPI network described by (von Mering et al., 2002) is henceforth denoted by ‘YHC’. This dataset is discussed in more detail in the next subsection. We denote by ‘Y11K’ the yeast PPI network defined by the top 11 000 interactions in the (von Mering et al., 2002) classification. ’YIC’ denotes the ‘core’ yeast PPI network from (Ito et al., 2000) Y2H study. We denote by ‘YIP’ the entire yeast PPI network from (Ito et al., 2000). ’YK’ is the yeast PPI network from (Krogan et al., 2006)
obtained by TAP and matrix-assisted laser desorption/ionization-time of flight mass spectrometry and liquid chromatography tandem mass spectrometry, ‘YM’ is the yeast PPI network from MIPS (Mewes et al., 2002). ‘FH’ is the high-confidence part of the fruitfly PPI network from (Giot et al., 2003) in which a two-hybrid-based protein-interaction map of the fly proteome has been presented.

‘WE’ is the entire worm PPI network published by Li et al., (2004), where more than 4000 interactions were identified from Y2H screens, and ‘WC’ denotes the ‘core’ part of the worm PPI network also from (Li et al., 2004). By ‘WS’ we denote the worm PPI network from (Zhong and Sternberg, 2006), where prediction techniques have been used to generate this PPI network, consisting of 18183 interactions.

The human PPI network from (Stelzl et al., 2005), obtained by Y2H screens, which contains high- medium- and low-confidence data is denoted by ‘HSL’, its part that contains only high- and medium-confidence data ‘HSM’, and only high-confidence interaction from this study by ‘HS’. ‘HR’ is the human PPI network from (Rual et al., 2005), also obtained by Y2H screens. By ‘HH’ we denote the human PPI network from the human protein reference database (HPRD) (Peri et al., 2004). We denote by ‘HM’ the human PPI network from MINT (Zanzoni et al., 2002).

5.2 Results

Using the algorithm described in section 4, we embedded these networks in 2D, 3D and 4D space. The resulting areas under the ROC curves are shown in Figure 3. One striking feature is that using only two dimensions typically gives results that are as good as the cases where dimension three or four is used. Of the 19 PPI networks, all but the YM 2D case produce an area under the ROC curve > 0.6, and 11 networks have areas under the ROC curves above 0.75. Note that some of the human PPI networks (Rual et al., 2005; Stelzl et al., 2005) that we analysed come from the first Y2H studies of the human interactome and thus are considered to be of low confidence (‘HR’, ‘HSL’, ‘HSM’, and ‘HS’ in Figure 3). We expect that low areas under the ROC curve for these networks are due to the noise present in them. Human PPI networks from curated databases MINT and HPRD are of higher confidence than the PPI networks from Y2H studies resulting in high areas under the ROC curve (see ‘HM’ and ‘HH’ in Figure 3).

5.3 Further tests

5.3.1 High-confidence data  Yeast S. Cerevisiae is an organism important for research in human biology. Von Mering et al., (2002) performed a systematic synthesis and evaluation of PPIs obtained using the main high-throughput PPI detection methods for yeast. They integrated 78 390 interactions between 5321 yeast proteins, out of which 2455 are identified by more than one PPI detection method (von Mering et al., 2002). This high-confidence PPI network, which has 2455 interactions amongst 988 proteins, appears as YHC in Figure 3. The actual areas under the ROC curve are 0.892, 0.893, 0.896 for embedding into 2D, 3D and 4D space, respectively. This represents a very good match to the geometric model and, reassuringly, it is the best over all the datasets. Hence, in our further investigations we will focus on this YHC data.

5.3.2 Geometric structure in other random network models  Modelling real-world networks by various types of random graphs began with the work of Erdős and Rényi. One classical random graph model connects nodes uniformly at random with some fixed probability (Erdős and Rényi, 1959, 1960). This simple model does not describe many important properties of real-world networks such as degree distribution and clustering coefficients. Efforts to improve the applicability of these networks produced generalized random graphs (Bender and Canfield, 1978) in which the edges are chosen at random as in Erdős–Rényi graphs, but the degree distribution matches the degree distribution of the real network. Attempts to further improve global properties of the real-world networks led to numerous types of models such as small-world (Watts and Strogatz, 1998) and scale-free (Barabási and Albert, 1999; Simon, 1955) networks. Other models have been constructed with the idea of simulating some biological and topological properties of real biological networks, e.g. stickiness model (Pržulj and Higham, 2006).

In Figure 4 we present the results of an experiment to measure the extent to which several random graph models can produce a geometric structure. Here, seven types of random networks have been generated corresponding to, i.e. having the same number of nodes and edges as, the dataset YHC: Erdős–Rényi random graphs (denoted by ‘ER’), random graphs with the same degree distribution as the data (denoted by ‘ER-DD’), 2D geometric random graphs (‘GEO-2D’), 3D geometric random graphs (‘GEO-3D’), 4D geometric random graphs (‘GEO-4D’), scale-free Barabási–Albert model graphs (‘SF’) and stickiness model graphs (‘STICKY’). Note that ‘ER-DD’, ‘SF’, and ‘STICKY’ are three different types of scale-free networks. For each type, 30 networks have been generated, embedded in 2D, 3D and 4D space, and the area under the ROC curve has been computed. Results are also shown for networks obtained from randomly rewiring 10% (‘G-10P’), 20% (‘G-20P’), and 30% (‘G-30P’) of edges in a 3D geometric
random graph, in an attempt to account for false positives and false negatives (further details in Section 5.3.3). The mean and the SD for AUC of these networks are shown in Figure 4, and we see that the algorithm clearly distinguishes between geometric and non-geometric models including scale-free networks. In other words, even allowing for noise in the data, the lack of geometric structure in the other models is apparent.

5.3.3 Robustness of our approach  Unfortunately, noise is inherent in all current PPI networks (Mrowka et al., 2001). On one hand, PPI datasets contain a large percentage of false positive interactions. One example is when proteins interact indirectly, i.e. through mediation of one or more molecules, but this is recorded as a direct physical interaction by the experimental method. On the other hand, imperfect experimental methods lead to false negative interactions. Different biochemical techniques produce different sets of false positives and negatives. Thus, trying to find high-confidence PPI networks by overlapping multiple datasets may result in discarding many real interactions.

Since the PPI networks are thought to contain a large percentage of false negatives and a large percentage of false positives, we tested with simulated noise. From Figure 4, we see that for randomly generated GEO-3D graphs the embedding into 3D space is excellent. At the right in Figure 4 we show the effect of rewiring 10%, 20% and 30% of the edges in these GEO-3D graphs. The resulting networks are denoted by G-10P, G-20P and G-30P, respectively. We generated 30 networks of each type (corresponding to the percentages of rewired edges), embedded them in 2D, 3D and 4D space, and computed the areas under the ROC curve. The mean area under the ROC curve for the 10% rewired GEO-3D networks are: 0.811, 0.875 and 0.918, corresponding to 2D, 3D and 4D embeddings, respectively.

In Figure 5, ROC curves corresponding to embeddings into 3D Euclidean space of the high-confidence yeast PPI network (denoted by YHC) and the model networks ER, ER-DD, GEO-3D and SF generated with the same number of nodes and edges as the YHC network, are presented in the same graph for comparison. Also, we included one GEO-3D network with 10% rewired edges (denoted by GEO-3D-10%). We see low areas under ROC curves for random (ER) and scale-free (ER-DD and SF) network types. We also see that the YHC–ROC curve is consistent with that of a noisy geometric network. So, overall, from the ROC curve perspective, any departure from geometric structure in the PPI network can be explained by the inherent noise.

6 DISCUSSION

It has already been established that the random geometric graph model gives an excellent fit for various global and local measures of PPI networks such as pathlengths, clustering coefficients, relative graphlet frequencies (Przulj et al., 2004), and graphlet degree distributions (Przulj, 2006). The main idea of this work is to test directly whether PPI networks are geometric by embedding them into a low-dimensional Euclidean space. We developed an algorithm that takes PPI network data and attempts to recover the geometric network structure, using specificity and sensitivity measures to quantify the results. The algorithm was demonstrated to work well on artificially constructed geometric random networks, even in the presence of noise. We applied the algorithm to the 19 PPI networks of various organisms (yeast, fruitfly, worm and human) and seven types of random network models including three types of scale-free networks. Also, we compared these results with the results of rewired geometric networks, where rewiring simulates the noise that is present in the real PPI network data. The results we obtained in this work yield support to the hypothesis that the structure of PPI networks is consistent with the structure of a noisy geometric random graph. The fact that the algorithm produced a better fit on high-confidence PPI data suggests that the algorithm could be
used to help discover false positives and false negatives in PPI networks.

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REFERENCES


